Academy of Mathematics and Systems Science

Chinese Academy of Sciences

Name: Xiayang Li(李夏洋) UID: 202328000206057 Mar. 28, 2024 Assignment

Stochastic Operations Research-Homework 1

Question 1

Both E and F are two events. If P(E) = 0.9 and P(F) = 0.8, show that $P(EF) \ge 0.7$. In general, show that

$$P(EF) \geqslant P(E) + P(F) - 1. \tag{1}$$

This is known as Bonferroni's inequality.

Solution. In fact, $P(EF) = P(E) + P(F) - P(E \cup F)$. As $P(E \cup F) \leq 1$, we obtain

$$P(EF) = P(E) + P(F) - P(E \cup F)$$

$$\geqslant P(E) + P(F) - 1.$$
 (2)

Question 2

Let X be a random variable, prove that $E[X^2] \ge (E[X])^2$. When do we have equality?

Solution. Note that

$$E[(X - E(X))^2] \geqslant 0, \tag{3}$$

and

$$E[(X - E(X))^{2}] = E(X^{2}) - E(X)^{2}.$$
(4)

Therefore,

$$E(X^2) \geqslant E(X)^2. \tag{5}$$

When and only when E(X) = X a.s., the equality holds.

Question 3

Let c be a constant and X be a random variable. Show that

- $Var(cX) = c^2 Var(X)$
- Var(c+X) = Var(X)

Solution. It is obvious that E(cX) = cE(X), $E(cX)^2 = cE(X)^2$, and E(X+c) = E(X). As $Var(X) = E(X^2) - E(X)^2$, the results follow.

Show that if X is nonnegative continuous random variable with distribution F, then

$$E[X] = \int_0^\infty (1 - F(x))dx. \tag{6}$$

For a general continuous random variable X with finite mean, use $X=X^+-X^-$ to prove

$$E[X] = \int_0^\infty (1 - F(x))dx - \int_{-\infty}^0 F(x)dx$$
 (7)

Solution. Denote that $\bar{F}(x) = 1 - F(x)$, we have

$$\bar{F}(x) = \int_{x}^{\infty} f(t)dt. \tag{8}$$

Therefore,

$$\int_{0}^{\infty} \bar{F}(x)dx = \int_{0}^{\infty} \int_{x}^{\infty} f(t)dtdx$$

$$= \int_{0}^{\infty} \int_{0}^{t} f(t)dxdt, \text{ (Xis nonnegative)}$$

$$= \int_{0}^{\infty} t f(t)dt$$

$$= E(X).$$
(9)

Besides, for a general continuous random variable X with the finite mean, use $X = X^+ - X^-$ and equation (9) we have

$$E[X] = E[X^{+}] - E[X^{-}]$$

$$= \int_{0}^{\infty} (1 - F_{X^{+}}(x))dx - \int_{0}^{\infty} (1 - F_{X^{-}}(x))dx.$$
(10)

Also, $1 - F_{X^+}(x) = 1 - F(x), x \ge 0$, and,

$$1 - F_{X^{-}}(x) = P(X^{-} > x) = P(X < -x) = F(-x), x \ge 0,$$
(11)

we obtain

$$E(X) = \int_0^\infty (1 - F(x)) dx - \int_0^\infty (1 - F_{X^-}(x)) dx$$

$$= \int_0^\infty (1 - F(x)) dx + \int_0^\infty (F(-x)) d(-x)$$

$$= \int_0^\infty (1 - F(x)) dx - \int_{-\infty}^0 (F(x)) dx.$$
(12)

Till now, we have completed this solution.

Question 5

Prove that if random variables X and Y are jointly continuous, then $\mathrm{E}[X] = \int_{-\infty}^{\infty} \mathrm{E}[X|Y = y] f_Y(y) dy$

Solution. We have

$$\int_{-\infty}^{\infty} E[X|Y=y] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} x f(x,y) dx dy$$

$$= \int_{-\infty}^{\infty} x \int_{-\infty}^{\infty} f(x,y) dy dx .$$

$$= \int_{-\infty}^{+\infty} x f_X(x) dx = E(X)$$
(13)

The solution has been completed.

Question 6

Let X_1 and X_2 be independent exponential random variables, each having rate μ . Let

$$X_{(1)} = \min\{X_1, X_2\} \text{ and } X_{(2)} = \max\{X_1, X_2\}.$$
 (14)

Find

- $E[X_{(1)}]$ and $Var(X_{(1)})$,
- $E[X_{(2)}]$ and $Var(X_{(2)})$

Solution. Note that

$$P(\min\{X_1, X_2\} > x) = P(X_1 > x)P(X_2 > x)$$

$$= e^{-\mu x} e^{-\mu x} , \qquad (15)$$

$$= e^{-2\mu x}.$$

we have

$$f_{X_{(1)}}(x) = 2\mu e^{-2\mu x} \sim \text{Exp}(2\mu).$$
 (16)

Therefore, $\mathrm{E}[X_{(1)}]=1/2\mu$ and $\mathrm{Var}[X_{(1)}]=1/(4\mu^2)$ follow.

Similarly,

$$P(\max\{X_1, X_2\} \leqslant x) = P(X_1 \leqslant x)P(X_2 \leqslant x)$$

$$= (1 - e^{-\mu x})(1 - e^{-\mu x})$$

$$= 1 - 2e^{-\mu x} + e^{-2\mu x}$$
(17)

We have

$$f_{X_{(2)}}(x) = 2\mu e^{-\mu x} - 2\mu e^{-2\mu x}.$$
 (18)

Then $\mathrm{E}[X_{(2)}] = \frac{3}{2\mu}$ and $\mathrm{Var}[X_{(2)}] = \frac{7}{4\mu^2}$ follow.

Question 7

Two individuals, A and B, both require kidney transplants. If she does not receive a new kidney, then A will die after an exponential time with rate μ_A , and B after an exponential time with rate μ_B . New kidneys arrive in accordance with a Poisson process having rate λ . It has been decided that the first kidney will go to A (or to B if B is alive and A is not at that time) and the next one to B (if still living).

- (a) What is the probability that A obtains a new kidney?
- (b) What is the probability that neither A nor B obtains a new kidney?
- (c) What is the probability that both A and B obtain new kidneys?

Solution. Let X_0 denote the arrived time of the first kidney, X_1 denote the time of death for A, and X_2 for B, and they are independent exponential random variables with rate λ , μ_A , μ_B .

Then the probability that the problem (a) is valid is equivalent to

$$P(X_0 < X_1) = \int_0^\infty P\{X_0 < X_1 | X_0 = x\} \lambda e^{-\lambda x} dx$$

$$= \frac{\lambda}{\lambda + \mu_A}$$
(19)

Also, the results for (b) is followed as

$$P(X_{0} < \max\{X_{1}, X_{2}\}) = \int_{0}^{\infty} P(X_{0} < \max\{X_{1}, X_{2}\} | X_{0} = x\} \lambda e^{-\lambda x} dx$$

$$= \int_{0}^{\infty} (e^{-\mu_{A}x} + e^{-\mu_{B}x} - e^{-(\mu_{A} + \mu_{B})x}) \lambda e^{-\lambda x} dx$$

$$= \frac{\lambda}{\lambda + \mu_{A}} + \frac{\lambda}{\lambda + \mu_{B}} - \frac{\lambda}{\lambda + \mu_{A} + \mu_{B}}.$$
(20)

The result for (c) is equivalent to

$$\begin{split} &P(N(X_1)=1,N(X_2)=2)\\ &=P(N(t)=1,N(X_2)=2|X_1=t)\\ &=\int_0^\infty P(N(t)=1|X_1=t)P(N(X_2-t)=1|X_1=t)dt\\ &=\int_0^\infty \int_0^\infty P(N(t)=1|X_1=t)P(N(s-t)=1|X_1=t,X_2=s)P(X_1=t)P(X_2=s)dsdt\\ &=\int_0^\infty \int_t^\infty P(N(t)=1|X_1=t)P(N(s-t)=1|X_1=t,X_2=s)P(X_1=t)P(X_2=s)dsdt \\ &=\int_0^\infty (1-e^{-\lambda t})\mu_A e^{-\mu_A t} \int_t^\infty (1-e^{-\lambda(s-t)})\mu_B e^{-\mu_B s}dsdt\\ &=\mu_A \mu_B \int_0^\infty (1-e^{-\lambda t})e^{-\mu_A t} (\frac{1}{\mu_B}-\frac{1}{\mu_B+\lambda})e^{-\mu_B t}dt\\ &=\frac{\lambda^2 \mu_A}{(\mu_A+\mu_B)(\mu_B+\lambda)(\mu_A+\mu_B+\lambda)} \end{split}$$

Let $\{N(t), t \ge 0\}$ be a Possion process with rate λ . For $i \le n$ and s < t,

- (a) find $P\{N(t) = n | N(s) = i\};$
- (b) find $P\{N(s) = i | N(t) = n\};$

Solution. Considering (a), we find that

$$P\{N(t) = n | N(s) = i\} = P\{N(t - s) = n - i\}$$

$$= e^{-\lambda(t - s)} \frac{(\lambda(t - s))^{n - i}}{(n - i)!}$$
(22)

Considering (b), we find that

$$P\{N(s) = i | N(t) = n\} = \frac{P\{N(s) = i, N(t) = n\}}{P\{N(t) = n\}}$$

$$= \frac{P\{N(s) = i\}P\{N(t - s) = n - i\}}{P\{N(t) = n\}}$$

$$= \binom{n}{i} \frac{s^{i}(t - s)^{n - i}}{t^{n}}.$$
(23)

The solution has been completed.

Question 9

Suppose that people immigrate into a territory according to a Poisson process with rate $\lambda = 2$ per day.

- (a) Find the probability these are exactly 12 arrivals in the following week(7 days).
- (b) Find the expected number of days until there have been 20 arrivals.
- (c) Given that there are 4 arrivals in a given two days, what is the probability that all 4 arrivals in the first one day.
- (d) Compute $E[S_4|N(2)=2]$

Solution.

(a)
$$P{N(7) = 12} = e^{-14\frac{14^{12}}{12!}}$$
.

(b)
$$E[S_{20}] = 20 \times \frac{1}{2} = 10.$$

(c)
$$P\{N(1) = 4 | N(2) = 4\} = \frac{P\{N(1) = 4, N(2) = 4\}}{P\{(N(2) = 4\}}$$

$$= \frac{P\{N(1) = 4, N(2) - N(1) = 0\}}{P\{N(2) = 4\}}$$

$$= \frac{P\{N(1) = 4\}P\{N(1) = 0\}}{P\{N(2) = 4\}}$$

$$= \frac{1}{16}$$
(d)

$$E[S_4|N(2) = 2] = 2 + E[S_2]$$

= 3

Question 10

If $\{N(t), t \ge 0\}$ is a renewal process and $S_n = \sum_{i=1}^n X_i$. Is it true that

- (a) N(t) < n if and only if $S_n > t$?
- (b) $N(t) \leq n$ if and only if $S_n \geq t$?
- (c) N(t) > n if and only if $S_n < t$?

Solution. (a) is true.

- (b) is not true. $S_n \ge t$ implies $N(t) \le n$ while the reverse is not true.
- (c) is not true. N(t) > n implies $S_n < t$. While $S_n < t$ can not imply N(t) > n.